Statistical Inference Models for Image Datasets with Systematic Variations

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Statistical analysis of longitudinal or cross sectional brain imaging data to identify effects of neurodegenerative diseases is a fundamental task in various studies in neuroscience. However, when there are systematic variations in the images due to parameter changes such as changes in the scanner protocol, hardware changes, or when combining data from multi-site studies, the statistical analysis becomes problematic. In this scenario, the goal is to develop a unified statistical solution to the problem of systematic variations in statistical image analysis. Based in part on recent literature in harmonic analysis on diffusion maps, we propose an algorithm which compares operators that are resilient to the systematic variations. These operators are derived from the empirical measurements of the image data and provide an efficient surrogate to capturing the actual changes across images. To evaluate the proposed ideas, we present various experimental results on detecting changes in simulations as well as how the method offers improved statistical power in the analysis of real longitudinal PIB-PET imaging data acquired from participants at risk for Alzheimer’s disease (AD).

High level description of the framework. Consider two unknown latent functions \( f \) and \( f' \), and let \( \alpha \) and \( \beta \) denote two parameters such that they modify the form of the function. Then, we are provided with the measurements, \( f_\alpha \) and \( f'_\beta \), i.e., both the parameter and the function changes. It is clearly not possible to verify the true difference between the latent functions, unless we also know the relationship between the transformations induced by \( \alpha \) and \( \beta \). Assume that an oracle provides us an operator \( T \) with the interesting property that it is invariant to the parameter space \( \mathcal{P} \) from which \( \alpha \) and \( \beta \) are drawn. That is, if we construct a pair of operators from the empirical measurements of \( f_\alpha \) and \( f'_\beta \), the operators will reflect the true difference between \( f \) and \( f' \). Since the operator \( T \) only offers invariances to the parameter space \( \mathcal{P} \), in this case, the operators \( T_{f_\alpha} \), and \( T_{f'_\beta} \) cannot be compared. Nonetheless, these operators provide a mapping to two different spaces, \( S_{f_\alpha} \) and \( S_{f'_\beta} \), since \( f \) and \( f' \) are distinct. Because of the invariance to \( \mathcal{P} \), if we plug a known function (such as an impulse function) at all locations in the original image into the two operators, we will obtain its transformed representations in \( S_{f_\alpha} \) and \( S_{f'_\beta} \). Once these transformed forms of the impulse functions are mapped from \( S_{f_\beta} \) to \( S_{f_\alpha} \), we can calculate the distance [1].

Key Idea. We assume that the two images \( f_\alpha \) and \( f'_\beta \) are spatially registered so that the only variations in the measurements comes from the parameters, \( \alpha, \beta \in \mathcal{P} \). We also assume that the relationship of measurements at each grid point to measurements at other grid points in the same image is preserved and is independent of changes in the parameter space, \( \mathcal{P} \). When we place a unit energy at location \( p \), the propagation of the energy will now show different patterns for the two image-derived operators: capturing the difference between those patterns is an excellent surrogate for detecting the difference between the two original functions \( f \) and \( f' \).

Interestingly, the process of constructing an operator and applying it on a \( \delta \) function is the same as the implementation of a mother wavelet function. In a graph setting (with \( N \) number of nodes), this mother wavelet is constructed using bases from spectral graph theory. Let \( \chi \) and \( \lambda \) denote the eigenvalues and eigenvectors from graph Laplacian, then the actual mother wavelet \( \psi_{\tau, p} \) at scale \( s \) centered at \( p \) is realized by a delta function \( \delta_p \) as

\[
\psi_{\tau, p}(q) = \mathcal{T}_s \delta_p = \sum_{l=0}^{N-1} g(s\lambda_l)\chi_l(p)\chi_l(q)
\]

where \( g() \) is a band-pass filter from a filter bank [2]. Here, a wavelet function \( \psi_{\tau, p}(q) \) can be viewed as a kernel function as \( \psi_{\tau, p}(q) \), defining a relationship between vertex \( p \) and \( q \). For our application, it is used to define Wavelet Kernel Distance (WKD) \( d_s(p, q) \) at scale \( s \), a measure between two points \( p \) and \( q \) defined as \( \ell^2 \)-norm of the wavelet density difference as

\[
d_s(p, q)^2 = \sum_{l=0}^{N-1} g(s\lambda_l)^2(\chi_l(p) - \chi_l(q))^2
\]

It can be interpreted as if we were comparing the effect of the same wavelet function dissipating from locations \( p \) and \( q \) to their neighbors by the wavelet kernel function \( g() \), thereby measuring the effect of the propagation.

Consider two individual graphs \( I \) and \( J \), constructed using functions (or images) \( f_\alpha \) and \( f'_\beta \) where the number of vertices in each is \( N \). Let \( \lambda^I, \lambda^J \) and \( \chi^I, \chi^J \) be the eigenvalues and eigenvectors from graph Laplacians of \( I \) and \( J \) respectively. On these graphs, WKD between a vertex \( p^I \) from \( I \) and a vertex \( q^J \) on \( J \) is defined as

\[
d_s(p^I, q^J)^2 = \sum_{l=0}^{N-1} g(s\lambda^J_l)^2\chi^J_l(p^I) + \sum_{l=0}^{N-1} g(s\lambda^J_l)^2\chi^J_l(q^J) - 2\sum_{l=0}^{N-1} g(s\lambda^J_l)^2\chi^J_l(p^I)\chi^J_l(q^J)
\]

using wavelet kernel functions \( \psi^I_{\tau, p} \) and \( \psi^J_{\tau, q} \).

Results. We applied our algorithm on synthetic images and for statistical analysis of longitudinal Pittsburgh compound B positron emission tomography (PIB-PET) scans. Representative results are shown in Fig. 1 and 2. In Fig. 1, we show a result comparing two different satellite images that cannot be compared directly due to image intensity transformation and in Fig. 2, we show correlation between a well-known AD risk factor (i.e., ratio of total \( \tau \)-protein and \( \beta \)-protein (1-42)) and subject age when a proper normalization is unavailable. More experimental results are provided in the main paper.
